| 1  | Title: Global warming may significantly increase childhood anemia burden in sub-  |
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| 2  | Saharan Africa  |
| 3  |   |
| 4  | Author list and affiliations:   |
| 5  | Yixiang Zhu <sup>1,†</sup> , Cheng He <sup>1, 2, 3†</sup> , Antonio Gasparrini <sup>4, 5, 6</sup> , Ana Maria Vicedo-Cabrera <sup>7,</sup>                      |
| 6  | <sup>8</sup> , Cong Liu <sup>1, 2</sup> , Jovine Bachwenkizi <sup>9</sup> , Lu Zhou <sup>1</sup> , Yuexin Cheng <sup>10</sup> , Lena Kan <sup>11</sup> , Renjie |
| 7  | Chen <sup>1, 2, *</sup> , Haidong Kan <sup>1, 2, 12, *</sup>  |
| 8  |   |
| 9  | <sup>1</sup> School of Public Health, Key Lab of Public Health Safety of the Ministry of Education  |
| 10 | and NHC Key Lab of Health Technology Assessment, Fudan University, Shanghai   |
| 11 | 200032, China   |
| 12 | <sup>2</sup> IRDR ICoE on Risk Interconnectivity and Governance on Weather/Climate Extremes   |
| 13 | Impact and Public Health, Fudan University, Shanghai 200438, China  |
| 14 | <sup>3</sup> Helmholtz Zentrum München -German Research Center for Environmental Health   |
| 15 | (GmbH), Institute of Epidemiology, Neuherberg, Germany  |
| 16 | <sup>4</sup> Department of Public Health, Environments and Society, London School of Hygiene &  |
| 17 | Tropical Medicine, London, United Kingdom   |
| 18 | <sup>5</sup> Centre for Statistical Methodology, London School of Hygiene and Tropical Medicine,  |
| 19 | London, UK  |
| 20 | <sup>6</sup> Centre on Climate Change and Planetary Health, London School of Hygiene and Tropical   |
| 21 | Medicine, London, UK  |
| 22 | <sup>7</sup> Institute of Social and Preventive Medicine, University of Bern, Bern, Switzerland   |
| 23 | <sup>8</sup> Oeschger Centre for Climate Change Research, University of Bern, Bern, Switzerland   |

| 24       | <sup>9</sup> Department of Environmental and Occupational Health, Muhimbili University of Health |
|----------|--|
| 25       | and Allied Sciences, Tanzania  |
| 26       | <sup>10</sup> Department of Hematology, The First People's Hospital of Yancheng, Yancheng        |
| 27       | Affiliated Hospital of Xuzhou Medical University, The Fourth Affiliated Hospital of              |
| 28       | Nantong University, Yancheng, China  |
| 29       | <sup>11</sup> Bloomberg School of Public Health, Johns Hopkins University, Baltimore, MD, United |
| 30       | States   |
| 31       | <sup>12</sup> Children's Hospital of Fudan University, National Center for Children's Health,    |
| 32       | Shanghai, China  |
| 33       |  |
| 34       | <sup>†</sup> Yixiang Zhu and Cheng He contributed equally to this work.                          |
| 35       |  |
| 36       | Corresponding author:  |
| 37       | *Dr. Renjie Chen, Department of Environmental Health, School of Public Health, Fudan             |
| 38       | University, P.O. Box 249, 130 Dong-An Road, Shanghai 200032, China. E-mail:                      |
| 39       | chenrenjie@fudan.edu.cn;   |
| 40       | **Dr. Haidong Kan, Department of Environmental Health, School of Public Health, Fudan            |
| 41       | University, P.O. Box 249, 130 Dong-An Road, Shanghai 200032, China. E-mail:                      |
| 42       | kanh@fudan.edu.cn.   |
| 43<br>44 | Lead Contact   |
| 45       | Dr. Haidong Kan, Department of Environmental Health, School of Public Health, Fudan              |
| 46       | University, P.O. Box 249, 130 Dong-An Road, Shanghai 200032, China. E-mail:                      |
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## 48 Summary

49 Childhood anemia constitutes a global public health problem, especially in low- and middle-income countries (LMICs). However, it remains unknown whether global warming 50 has an impact on childhood anemia. Here, we examined the association between annual 51 52 temperatures and childhood anemia prevalence in sub-Saharan Africa, and then projected childhood anemia burden attributable to climate change. Each 1°C increment in annual 53 temperature was associated with increased odds of childhood anemia (odd ratio=1.138, 95% 54 confidence interval: 1.134–1.142). Compared with the baseline period (1985–2014), the 55 attributable childhood anemia cases would increase by 7,597 per 100,000 person-year 56 under high emission scenario in 2090s, which would be almost 2-fold and over 3-fold 57 more than those projected in moderate and low emission scenarios. Our results reveal the 58 vulnerabilities and inequalities of children for the excess burden of anemia due to climate 59 60 warming, and highlight the importance of climate mitigation and adaptation strategies in LMICs. 61

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Keywords: Climate change; high temperature; anemia; disease burden; sub-Saharan
Africa

## 65 Introduction

66 Childhood anemia constitutes a global public health problem, especially in low- and 67 middle-income countries (LMICs).<sup>1</sup> Based on estimates from the World Health 68 Organization (WHO), 269.4 million children under 5 years suffered from anemia globally 69 in 2020 and 38.0 % of the cases were from sub-Saharan Africa.<sup>2,3</sup> Children with anemia 70 are at high risks of severe adverse health outcomes including cognitive or behavioral 71 impairment,<sup>4</sup> poor growth and development,<sup>5</sup> and even death.<sup>6</sup>

The causes of childhood anemia are multifactorial and interrelated in complex ways. 72 The etiology of anemia could be related to proximate determinants (e.g., food insecurity, 73 contaminated water and sanitation problem), as well as some immediate causes, including 74 nutritional deficiencies (iron, vitamins A and B<sub>12</sub>, etc.), infection (e.g., soil-transmitted 75 helminth infection, schistosomiasis, and malaria), systemic inflammation, and genetic 76 hemoglobin (Hb) disorders.<sup>7</sup> In addition, epidemiological evidence has indicated that 77 environmental factors, such as high ambient temperature, may have some relationships 78 with hemoglobin concentration and anemia status in children.<sup>8</sup> However, the possible roles 79 of climate factors have rarely been considered in designing, monitoring, prevention and 80 control programs of anemia.<sup>9,10</sup> 81

Climate change has become the largest global threat for human health in the 21st century.<sup>11</sup> Compared to the recent decades (1995–2014), global surface air temperature is likely to increase by 2.4°C–4.8°C by the end of the 21st century with the assumption of unrestricted carbon emission, beyond the 2°C global temperature target of the Paris Agreement.<sup>12</sup> Sub-Saharan Africa is predicted to warm faster than the global average during the 21st century.<sup>13</sup> In addition, climate change is expected to disproportionately

affect vulnerable groups (such as children under 5 years), and further exacerbates existing 88 health inequities.<sup>14</sup> LMICs, which produce the least greenhouse gases emissions, have far 89 less capability to adapt to climate change.<sup>15</sup> Consequently, vulnerable populations in 90 LMICs will suffer from more health threats than those in high-income countries.<sup>16</sup> Climate 91 change is expected to reduce agricultural production,<sup>17</sup> alter micronutrient content of food 92 crops,<sup>18</sup> and cause food insecurity. On the other hand, high ambient temperatures may 93 shorten the breeding and development of mosquito, as well as the incubation period for the 94 Plasmodium parasites,<sup>19</sup> which could increase the transmission of malaria.<sup>20</sup> Malaria 95 parasites could infect red blood cells and cause them to rupture (a process known as 96 hemolysis), resulting in the spread of infectious anemia in children.<sup>21</sup> Warming climate 97 could thus lead to nutritional iron deficiency and susceptibility of parasitic infections, 98 contributing to the development of anemia in children. However, few studies have directly 99 linked ambient temperature with childhood anemia, and it remains poorly understood how 100 the burden of childhood anemia will vary in a warming climate.<sup>22</sup> 101

To understand the impacts of global warming on childhood anemia in LMICs, we first 102 explored the association of ambient temperature with the prevalence of childhood anemia 103 across sub-Saharan Africa, and then projected excess cases of childhood anemia 104 attributable to global warming under selected climate change scenarios by the end of 21st 105 century. We found each 1°C increment in annual temperature was associated with 106 increased odds of anemia (odd ratio=1.138, 95% confidence interval: 1.134-1.142) in 107 108 children. Compared with the baseline period (1985–2014), the excess childhood anemia cases attributable to climate change would increase by 7,597 (95% CI: 4,133–10,894) per 109 100,000 person-year under high emission scenario in 2090s, which would be almost 2-fold 110

and over 3–fold more than those projected in moderate and low emission scenarios (Table
S11). The results reveal the vulnerabilities and inequalities of children for the excess
burden of anemia due to climate warming, and highlight the importance of climate
mitigation and adaptation strategies in low- and middle-income countries.

115

### 116 **Results**

## 117 **Descriptive data**

Between 2003 and 2020, this study includes a total of 275,377 children under five 118 119 years old from 26 sub-Saharan African countries. The overall prevalence of childhood anemia is 63.76% (Table 1). Of anemia in different degrees, the prevalence of moderate 120 121 anemia is highest (35.34%), followed by mild anemia (25.06%) and severe anemia (3.47%) 122 (Table S1). Spatial distribution of childhood anemic cases and annual mean temperature 123 levels across the sub-Saharan Africa are shown in Figure 1. Specifically, anemia 124 prevalence is higher in Western Africa, Central Africa, and several subregions in Eastern Africa, where higher annual temperatures are also recorded. Southern Africa has a lower 125 126 anemia prevalence and lower temperature. In the country level, annual average 127 temperatures range from 16.3 °C in Lesotho to 28.8 °C in Burkina Faso (Table 1). Characteristics of the study population are summarized in Table S2. The mean age of 128 children was 2.11 (±1.33) years and 49.6% (N=136,575) were girls. 129

# 130 Baseline temperature-anemia association

Figure 2 shows the odds ratio (OR) estimates for the associations between annual mean temperature and childhood anemia prevalence at the region and country levels. The OR of total anemia is 1.138 [95% confidence interval (CI): 1.134–1.142] associated with

each 1°C increment in annual mean temperature. Regionally, there are higher ORs in 134 Eastern Africa, Central Africa and Western Africa than in Southern Africa. For specific 135 countries, Gabon has the largest OR, while the estimates in most Southern African 136 countries are not statistically significant. Specifically, for a 1°C increment in annual mean 137 temperature, increase in the prevalence of severe anemia (OR=1.259, 95% CI: 1.245–1.273) 138 139 is largest, followed by moderate anemia (OR=1.173, 95% CI: 1.169–1.178) and mild anemia (OR=1.090, 95% CI: 1.085–1.095) (Table S3). Table S4 presents the corresponding 140 risk ratio (RR) estimates in each country. The ORs of childhood anemia are larger when 141 longer averaging exposure period was used from 1 month to 12 months preceding the 142 interview day (Table S5). In addition, the main model exhibited better model fit compared 143 to the models in sensitivity analyses using shorter exposure periods (Table S6). 144

We further examined the possible mediation effects of childhood malnutrition and malaria infection. 11.40% and 9.74% of the association between annual mean temperature and childhood anemia are estimated to be mediated by the prevalence of childhood malnutrition and malaria infection, respectively (Table S7).

In Figure S1, the shape of exposure-response (E–R) relationship curve between annual mean temperatures and odds ratios of childhood anemia in sub-Saharan Africa is approximately linear within most of the temperature range; and the OR of childhood anemia consistently increases with higher annual temperature. There is no significant difference between linear and nonlinear models, supporting the empirically linear assumption.

In the sensitivity analysis, the main risk estimates are robust to the alternative removal
of household factors (wealth level, floor material, roof material, water infrastructures and

157 sanitation infrastructures), country-level covariates, or annual cumulative precipitation 158 (Table S8). When using a smaller lookback window (the life-time exposure) for the 159 children under 1-year-old, the results are also robust. The regional effect estimates are 160 robust to the use of random effect meta-analysis based on the country-level ORs (Table 161 S9).

# 162 **Temporal and spatial trends in temperature change**

Figure 3 shows the trends of projected temperature change from the baseline year by 163 region between 2015 and 2099 under three climate change scenarios (SSP1-2.6, SSP2-4.5, 164 165 and SSP5-8.5). We project a steep increase in annual mean temperatures across this century under the high-emission scenario (SSP5-8.5), while the increasing trends are modest under 166 the SSP2-4.5 and SSP1-2.6 scenarios which level off after the mid-21st century (Figure 167 **3A**). The annual temperatures show consistently increasing trends for all regions of sub-168 Saharan Africa, with the most prominent warming in Southern Africa (Figure 3B). For 169 different climate changing scenarios, the average annual temperatures in sub-Saharan 170 Africa would increase by 4.0°C (95% CI: 3.1°C–4.6°C), 2.2°C (95% CI: 1.7°C–2.5°C) and 171 1.1°C (95% CI: 0.8°C-1.5°C) for the SSP5-8.5, SSP2-4.5, and SSP1-2.6 scenarios, 172 173 respectively (Table S10). Regionally, Western Africa and Central Africa would experience higher temperature levels in the 2090s, whereas ambient temperature would increase more 174 175 prominently in Southern Africa by the end of 21st century.

176 **Projected anemia burden due to climate change** 

Figure 4 illustrates excess cases of childhood anemia per 100,000 person-year attributable to climate change in sub-Saharan Africa under three climate change scenarios (SSP1-2.6, SSP2-4.5, and SSP5-8.5). The spatial distributions of the projected disease

burden are similar for the three scenarios, but the magnitude of future changes in burden 180 differ considerably (Figure 4A). Under the assumption of no changes in population, anemia 181 patterns or climate adaption, there are increasing excess cases of childhood anemia in the 182 projected period (2015–2099) in sub-Saharan Africa, compared with the baseline (1985– 183 2014) (Figure 4B). The increasing trends are steeper under the SSP5-8.5 scenario, whereas 184 185 the trends are modest under SSP1-2.6 and become flat after the middle of this century under SSP2-4.5. Compared with the baseline period (1985–2014), the excess childhood anemia 186 cases attributable to climate change would increase by 7,597 (95% CI: 4,133–10,894) per 187 100,000 person-year under the SSP5-8.5 scenario in 2090s, which would be almost 2-fold 188 and over 3-fold more than those in SSP2-4.5 (4,208, 95% CI: 2,311-6,015) and SSP1-2.6 189 (2,275, 95% CI: 1,277–3,226) scenarios (Table S11). 190

Figure 4 also reveals different increasing patterns across regions. The excess cases of childhood anemia would rise more prominently in Central Africa than in other regions under SSP5-8.5, whereas there are nonsignificant increments in Southern Africa (Figure 4A). Taking the SSP5-8.5 scenario for an example, there would be increments of 10,566 (95% CI: 9,062–12,045) cases per 100,000 person-year in Central Africa, followed by 9,276 (6,637–11,854) cases in Eastern Africa and 6,646 (2,485–10,619) cases in Western Africa by the end of the 21st century (Table S11).

When grided future population of children under 5 years is included in the projection of childhood anemia burden, the changes of excess cases of childhood anemia related to climate warming are not significantly different from those derived from our main analyses under the assumption of no population change (see Table S11 and Table S12).

202

# 203 Discussion

This epidemiological investigation in 26 sub-Saharan African countries demonstrates 204 increased risk of childhood anemia prevalence associated with higher ambient temperature. 205 We found childhood malnutrition and malaria infection could mediate an appreciable 206 portion of the association between high annual mean temperature and childhood anemia, 207 208 respectively. Our study provides the first projection on the future burden of childhood anemia posed by global warming in sub-Saharan Africa, which would increase 209 significantly, especially under the SSP5-8.5 scenario. We also observed significant 210 regional disparities for the impacts of global warming on childhood anemia, with the 211 highest burden in Central Africa, followed by Western Africa and Eastern Africa. 212

Although no epidemiological studies have directly linked childhood anemia with 213 ambient temperature, there are some indirect population-based evidences supporting our 214 findings. A previous cross-sectional study found that land surface temperature was 215 inversely associated with mean hemoglobin concentration in preschool-age children, a core 216 biomarker of childhood anemia.<sup>8</sup> However, another study conducted in four sub-Saharan 217 African countries did not observe a significant association.<sup>23</sup> We observed significant 218 219 associations between warmer ambient temperatures and higher prevalence of childhood anemia in three sub-Saharan African regions (Eastern, Central and Western Africa), 220 wherein children in Central Africa suffered highest temperature-related anemia risk. 221 222 Previous studies indicated that Central Africa had less precipitation and more vulnerability to drought than other regions.<sup>24,25</sup> High temperature and low precipitation may co-223 exacerbate the childhood anemia condition in drought-prone regions.<sup>26,27</sup> On the other hand, 224 225 a nonsignificant association was observed in Southern African counties where annual

temperature was appreciably lower than other regions. The exposure-relationship curve in this study supports that the regions of lower annual temperature range will experience lower childhood anemia risk.

We calculated the excess cases of childhood anemia attributable to annual 229 temperatures. The projected burden of childhood anemia would increase in the context of 230 231 global warming, but the magnitude would be highly affected by the extent and timing of warming under different climate change scenarios. There is a steep increase in future 232 temperature, accompanied by a drastic increase in excess cases of childhood anemia under 233 SSP5-8.5, a scenario characterized by unrestricted greenhouse gas emissions, population 234 growth, energy consumption, and excessive land use. In contrast, the burdens of childhood 235 anemia would increase modestly or plateau in the middle of this century under SSP2-4.5 236 or stricter SSP1-2.6, which are in accordance with the respective trends of temperature 237 change in two scenarios. The findings emphasize the importance of developing targeted 238 climate mitigation strategies and adaptation strategies for preventing the exacerbation of 239 childhood anemia status in a changing climate. 240

The impact of climate change on childhood anemia seems to be biologically plausible. 241 242 Two major culprits of anemia, namely, malnutrition and parasitic infections, are both impacted under global warming. Previous studies have suggested that children's nutritional 243 244 status and susceptibility to pathogen infections may be affected by climate change through 245 several pathways. First, vulnerability of agriculture to climate change increases the risk of child malnutrition. The combination of rising temperature and increasing atmospheric 246 247 carbon dioxide may lead to child malnutrition and micronutrient deficits (such as iron deficiency) by attenuating agricultural productivity and crops nutrient content.<sup>28,29</sup> Second, 248

heat stress could reduce the appetite and further result in malnutrition because of 249 physiological adaption to high temperatures.<sup>30</sup> Third, higher temperature will promote the 250 transmission of parasites and increase the risk of malaria infection among vulnerable 251 populations, consequently increasing the malaria-related anemia in warmer climate.<sup>31</sup> 252 Furthermore, our analysis indicated that childhood malnutrition and malaria infection 253 254 would partly mediate the impacts of high ambient temperature on childhood anemia, strengthening the biological plausibility of our epidemiological findings.<sup>32-35</sup> High 255 temperature exposure could also induce a series of pro-inflammatory responses,<sup>36</sup> and thus 256 disturb iron metabolism and red blood cell production, potentially increasing the risk of 257 childhood anemia.<sup>37,38</sup> In addition, high temperatures accompanied by drought can also 258 lead to water scarcity, and the resulting poor water, sanitation and hygiene conditions may 259 lead to environmental enteric dysfunction, an underlying cause of anemia in children.<sup>39,40</sup> 260 We applied different lengths of exposure periods in exploring the association between 261 ambient temperature and prevalence of childhood anemia, and found that the association 262 became stronger when longer exposure period was used. The short-term impacts of the 263 flood on infant mortality could be interpreted by several pathways. According to previous 264 studies, relative shorter-term high temperature exposure (1 to 3 months) may be linked with 265 infectious diseases which mainly contribute to infective anemia,<sup>41</sup> whereas longer-term 266 exposure (e.g., annual temperature) may be linked with water resource, agricultural 267 productivity and malnutrition, and consequently cause nutritional anemia.<sup>42</sup> 268

Although no previous studies predicted the childhood anemia burden in the context of global warming, our findings were also supported by previous projections on parasitic infections and malnutrition burden due to future climate change.<sup>31,43,44</sup> For example,

climate change would have a considerable impact on rates of severe stunting, which was 272 estimated to increase by 31%, 36% and 55% in Central, Western and Eastern Africa in 273 2050, respectively.<sup>45</sup> Similarly, Jesse et al reported a 37% increase in the prevalence of 274 child wasting associated with high temperature in Western Africa, while climate change 275 was predicted to result in a 25% increase in the proportion of children with wasting in 276 Eastern Africa and Central Africa by 2100 under the high-emission scenerio.<sup>44</sup> Although 277 heterogeneity exists in study period, high risk region and disease burden estimates between 278 our projections and previous studies, we found a similar temporal trend of increasing 279 burden of these diseases for children in sub-Saharan Africa. Moreover, epidemic belt 280 expansion and increased population at risk for malaria were reported under multiple 281 climate change sceneries. The population at risk of malaria would increase up to about 736 282 million additional people by 2070 in the scenario of RCP8.5-SSP2, which may also lead to 283 an increased burden of malaria-related anemia.<sup>31</sup> However, temperature-related burdens of 284 childhood anemia need to be fully captured in broader LMICs with unified analytical 285 frameworks. 286

Our study findings have notable public health implications. First, we projected that 287 288 temperature-related excess cases of childhood anemia would significantly increase under future climate change, especially under the high emission scenario. Considering the 289 challenges posed by global warming, we urgently need multi-sectoral collaborations to 290 291 address long-term policies and plans to reduce anthropogenic carbon emissions to slow down the pace of global warming. Second, climate change disproportionally affects the 292 293 populations in LMICs and vulnerable groups, adding threats to environmental equity and justice.<sup>16</sup> Their vulnerabilities to climate change stresses the urgency to implement 294

measures to mitigate climate change hazards, including the establishment of high 295 temperature alerting systems,<sup>46</sup> the effective allocation of medical resources,<sup>47</sup> etc. Third, 296 our results further suggested that pre-school aged children in LMICs suffered unevenly 297 from the impacts of climate change on anemia.<sup>48</sup> Local governments need to develop 298 sophisticated protective strategies for this vulnerable group, for example, increasing 299 popularization of air conditioning and indoor ventilation,<sup>49</sup> improving housing conditions 300 with thermal insulation materials,<sup>33</sup> and advocating the use of mosquito nets and iron 301 supplementation in high-temperature regions. <sup>34,35</sup> 302

303 Several limitations of this study should be acknowledged. First, the exposure-response associations between annual temperature and childhood anemia were obtained through a 304 cross-sectional study design, which attenuated the causality for our findings. Second, we 305 projected childhood anemia burden under the assumption of no changes in anemia patterns, 306 and climate adaptation, so our prediction may represent the upper limit of the contribution 307 of climate warming to anemia prevalence.<sup>50</sup> Third, our dataset only included LMICs in 308 tropical and sub-tropical regions, so caution should be taken when extrapolating our results 309 to other regions. 310

311

# 312 Conclusions

This multi-country study in sub-Saharan Africa demonstrates that higher environmental temperature could increase the prevalence of childhood anemia. Global warming would further significantly increase the burden of childhood anemia in this area. These results reveal the vulnerabilities and inequalities of children in LMICs in suffering from anemia under global warming, and highlight the importance of mitigation and adaptation strategies of climate change, especially for vulnerable subgroups in climate-

319 sensitive regions.

320

## 321 **Experimental Procedures**

## 322 **Resource Availability**

323 Lead contact

324 Requests for further information and resources should be directed to the lead contact,

325 Haidong Kan (kanh@fudan.edu.cn.).

# 326 Materials availability

327 No materials were used in this study.

## 328 Data and code availability

The data and code used in this analysis can be accessed online. Databases: (1) Data of 329 Demographic and Health Surveys (DHS) conducted in sub-Saharan Africa 330 (https://dhsprogram.com/); (2) Historical data of ambient temperature and precipitation 331 (https://cds.climate.copernicus.eu/); (3) Country-specific population data 332 (https://data.unicef.org/dv index/) and anemia prevalence (https://www.who.int/data/gho/) 333 334 of under 5-year-old children; (4) Country-specific Crop Production Index and Gross Domestic Product per capita (https://data.worldbank.org/indicator/); (5) Future 335 336 temperature exposure under different scenarios (https://esgf-node.llnl.gov/search/cmip6/); 337 (6) Future population size data (https://sedac.ciesin.columbia.edu/data/collection/gpw-v4) and the proportion of children under 5 years (https://tntcat.iiasa.ac.at/SspDb/) under 338 different scenarios. Code: source code for temperature-anemia association and future 339 340 childhood anemia burden projection (https://github.com/YixiangZhu/Anemiatemperature/). Any additional information required to reanalyze the data reported in thispaper are available from the lead contact upon request.

### 343 **Data sources**

The DHS program collected health, behavior and sociodemographic data routinely 344 (about every 5 years) in more than 90 countries, covering a series of topics such as maternal 345 346 and child health, malaria, domestic violence, environmental health.1 In DHS survey, a "cluster" is a group of adjacent households which serves as the primary sampling unit in 347 sampling procedure. Using a stratified two-stage cluster sampling design, clusters are 348 randomly selected from the areas stratified by geographic region and by urban/rural area 349 within each region, while the households are randomly selected from each cluster. From 350 eligible households, all ever-married women of reproductive age 15-49 were interviewed 351 by trained fieldwork staff, and data of children under 5 years were collected from all 352 clusters of the countries. 353

We included the following variables of children: age (<2 and 2-5 years), gender (male 354 and female), body mass index (BMI), result of malaria with rapid diagnostic test (RDT). 355 Mother's age (15-19, 20-29, 30-39, 40-49 years) and education level (no education, primary, 356 357 secondary, higher) are adjusted for as indicators of socio-economic status. We further controlled household characteristics including: type of residence (urban and rural), one or 358 359 more children in the household (singleton and multiple), type of water infrastructures 360 (surface water, well, piped/tap and others); type of sanitation infrastructures (no facility, latrine, flush toilet and others), materials of floor (natural, rudimentary, finished, others), 361 362 materials of roof (natural, rudimentary, finished, others) and wealth level. We calculated 363 the scores of household wealth index using a linear principal component analysis (PCA)

based on the data of household assets including electricity, car, television, refrigerator,
bicycle, scooter, mobile telephone and cooking fuel (Results of PCA were shown in Table
S13). Each survey separates interviewed households into five wealth quintiles (poorest,
poorer, middle, richer, richest) to characterize their wealth level.

## 368 Health outcome

369 In all surveys, children aged 5 years or younger in the eligible households were tested for anemia. Hemoglobin (Hb) concentrations in children were measured by finger- or heel-370 prick blood specimens using a portable Hemo Cue autoanalyzer.<sup>51</sup> In accordance with the 371 372 WHO definition of anemia in children under 5 years, a child was considered as anemic if his or her altitude-adjusted Hb level was less than 11 g/dL. An Hb levels below 7.0 g/dL 373 was classified as severe anemia, a level between 7.1g/dL and 9.9g/dL was classified as 374 moderate anemia and a level between 10.0 g/dL and 10.9 g/dL was classified as mild 375 anemia.<sup>52</sup> Informed consent to participate in DHS interviews and biomarker tests was 376 obtained orally from parents or guardians. 377

# 378 Historical meteorological parameters and population data

Historical data of ambient temperature and precipitation were derived from ERA-5, a 379 380 global reanalysis dataset on latitude–longitude grids at a resolution of approximately 0.25° ×0.25° and up to 1-h frequency, which was produced by the European Centre for Medium 381 Range Weather Forecasts (ECMWF).<sup>53</sup> We assigned daily mean temperature and daily 382 cumulative precipitation based on available geocoded coordinates for their clusters from 383 Jan 1, 2003, to Dec 31, 2020. Then, we computed the annual average temperature and 384 annual cumulative precipitation for each child under 5 years over the 365 days (one year) 385 prior to the interview day. 386

We obtained country-specific population data of under 5-year-old children for the survey year from the United Nations International Children's Emergency Fund (UNICEF) Data Warehouse. Prevalence of anemia in children aged 6–59 months in the corresponding country and survey year was obtained from the Global Health Observatory project of WHO. The Crop Production Index and Gross Domestic Product per capita for each sub-Saharan country in the survey year were derived from the World Bank.

393

# **Future meteorological parameters**

Future time series data of daily mean temperature for various climate change scenarios 394 were derived from the latest internationally-coordinated Coupled Model Intercomparison 395 Project sixth phase (CMIP6).<sup>12</sup> The projections of global climate change were assessed 396 based on the Shared Socioeconomic Pathways (SSPs), which include five common 397 scenarios (i.e., SSP1-1.9, SSP1-2.6, SSP2-4.5, SSP3-7.0, and SSP5-8.5). These scenarios 398 correspond to the increasing trajectories of atmospheric greenhouse gas concentration, and 399 describe a range of warming in global climate from mild (SSP1-1.9) to extreme (SSP5-8.5). 400 We selected three most common scenarios and the baseline period (1995-2014) in 401 accordance with previous projection studies.<sup>12,54</sup> Compared to the baseline period, global 402 surface air temperature is likely to increase 2.4°C–4.8°C in the high-emission scenario 403 (SSP5-8.5) over the period 2081–2100, and by  $0.5^{\circ}$ C–1.5°C and  $1.2^{\circ}$ C–2.6°C for the low 404 and moderate emission scenarios (SSP1-2.6 and SSP2-4.5), respectively.<sup>12</sup> Finally, we 405 406 extracted daily temperature data during the baseline period and the projection period (2015-2099) from 20 global climate models (GCMs) datasets for historical and future 407 408 temperature simulation in various climate change scenarios (Table S14).

Temperature data of GCM outputs were interpolated statistically to a geographical 409 grid of a  $1.0^{\circ} \times 1.0^{\circ}$  resolution using a bilinear interpolation method, and were then 410 transformed into city-level estimates by spatially averaging the gridded data within each 411 sub-Saharan African country. Cities and the boundaries were defined based on the 412 Database of Global Administrative Areas Version 4.1 (https://gadm.org/).<sup>55</sup> We 413 downloaded data of the included 26 countries and extracted the boundaries at 414 Administrative Level 2, which was deemed as the city level in this analysis. However, the 415 projected temperature series derived from different GCMs may result in nonnegligible bias 416 when data was applied to fitting the association between temperature and health outcomes 417 quantified by ERA-5 reanalysis data. Therefore, we extracted the modelled daily 418 temperature series for each grid in the studied country during 1985-2099, and further 419 corrected the modeled temperatures with data in corresponding grids from ERA-5 420 reanalysis series by using an additive scaling method for all GCMs.<sup>56</sup> 421

422 Future population data

We obtained predicted grided population size data under the SSP1 (Sustainability), 423 SSP2 (Middle of the Road) and SSP5 (Fossil-fueled development) scenarios from National 424 425 Aeronautics and Space Administration (NASA) Socioeconomic Data and Applications Center (SEDAC).<sup>57</sup> The database provides global urban, rural, and total population base 426 427 year and projection grids at a resolution of 1-km (about 30 arc-seconds) based on the SSPs. 428 We further obtained projected total population and population for children aged under 5 years at the country level, using the SSP Database - Version 2.0, under the same scenarios 429 above.<sup>58</sup> Due to the unavailability of population age structure projections for all countries, 430

we applied country-level projections to each location in 10-year intervals from 2020 to2100.

We then calculated the proportion of children under 5 years in a future year during 2020 to 2100 by dividing the projected age group-specific population for that year by the total population projected for the same year. Finally, we utilized the proportion of children under 5 years at the country level and applied it to the predicted population grid for the corresponding year and SSP, in order to acquire the population grid of children under 5 years for the future.

439 Statistical Analyses

# 440 **Temperature-anemia association**

Because previous studies have found associations between higher temperatures and 441 increases in common anemia risk factors, such as malnutrition and malaria infection,<sup>59,60</sup> 442 we hypothesized a linear E-R relationship between annual mean temperature and the 443 444 prevalence of childhood anemia, and only explored the impact of temperature increase in sub-Saharan Africa. Then, mixed effect multivariable logistic regression models were 445 applied to estimate the associations between annual mean temperature and anemic status 446 447 (binary variable) of children at individual level. Based on literatures about the risk factor of childhood anemia,<sup>61,62</sup> we adjusted for the following covariates in the main model: age, 448 449 gender, BMI, insecticide-treated net use, the type of residence, one or more children in the 450 household, mother's age and education level, household wealth level. We controlled for the type of WASH (water and sanitation, and hygiene) infrastructures and annual cumulative 451 452 precipitation, which may affect the spread of infectious diseases in population. Materials 453 of roof and floor were also adjusted, which could modify the effect of ambient temperature on the residents inside the house. To better isolate the effect of temperature increase, we control for decadal mean temperature and random effects for city in the main model. We also included other country-level covariates, including Crop Product Index and Gross Domestic Product per capita in the model. Finally, we controlled survey month and year to adjust for seasonality and long-term trends, respectively. The effect of ambient temperature on anemia status was presented as OR and its 95% CIs of childhood anemia prevalence associated with a 1°C increment in mean temperature over the past year.

461 To allow for the calculation of excess cases, we applied the following formula [1] to 462 estimate the relative risks (RRs) (and 95% CI) in each sub-Saharan African country.<sup>63</sup>

463 
$$RR_c = \frac{OR_c}{1 - Rate_c + Rate_c \times OR_c}$$
[1]

464 where  $RR_c$  is the estimated RR (and 95% CI) of childhood anemia incidence associated 465 with a 1°C increment of annual mean temperature in each country;  $Rate_c$  is the country-466 specific prevalence of childhood anemia over the corresponding survey year.

In addition, we examined whether the effect estimates vary by different exposure periods. Specifically, we calculated the average temperatures during the past 1 month, 3 months, 6 months and 9 months preceding the interview day, and then re-performed the main analysis for the association between temperature and childhood anemia prevalence in each exposure period.

Furthermore, we conducted a causal mediation analysis to examine the possible roles of childhood malnutrition and malaria infection on the association between high temperature exposure and childhood anemia. Childhood malnutrition was defined as a Kaup index (BMI in children) < 15 kg/m-2.<sup>64</sup> Data on childhood malaria infection was collected by RDT in DHS survey. Specifically, we fit the main model and the adjusted model (i.e., the main model adjusted by the presence of malnutrition or malaria infection) to calculate total effect (TE) and direct effect (DE), respectively. The indirect effect (IE) of the mediator was calculated by the association coefficient of high temperature exposure in the main model (TE) minus the corresponding coefficient in the mediator-adjusted model (DE). Then, the percentage of mediation effect was calculated as the proportion of IE in the TE from the main model. The empirical 95% CIs for IE was estimated using a bootstrap resampling with 1000 samples.

To examine the shape of temperature-anemia association and test non-linearity of the curve, we flexibly depicted the E–R relationship curve between annual mean temperature and childhood anemia prevalence using the generalized additive model with a natural cubic spline of 3 degrees of freedom for temperature and covariates. We used F tests in analysis of variance (ANOVA) to test difference between linear model and non-linear model.

We tested the robustness of temperature-anemia association in three sensitivity 489 490 analyses. First, we fit three separate models based on the main model: 1) leaving out the household factors (wealth level, floor material, roof material, water infrastructures and 491 sanitation infrastructures); 2) removing country-level covariates; and 3) leaving out annual 492 493 cumulative precipitation. Second, we conducted a random effect meta-analysis based on the country-level ORs to derive the regional effect estimates. Third, we conducted a 494 495 sensitivity analysis by adding gridded future population of children under 5 years in the 496 projection of anemia burden prevalence.

# 497 Future changes in anemia burden due to climate warming

To reveal the potential impact of climate change on anemia, we projected changes in childhood anemia burden due to warming temperature, assuming no changes of population, anemia patterns or climate adaption over the projection period (2015–2099). Under the assumption of linear temperature-anemia association, we calculated the historical and future numbers of excess cases of childhood anemia due to annual-mean temperature increase as:

504

$$Excess \ cases = Pop_c \times Rate_c \times ERC_c \times \Delta T$$
<sup>[2]</sup>

where  $Pop_c$  is the size of population under 5 years old at city-level in each country; 505 *Rate* represents the baseline annual prevalence of childhood anemia in each country during 506 the survey year; and ERCc represents the country-specific percentage change in the risk of 507 childhood anemia for a 1°C increment in temperature;  $\Delta T$  indicates the predicted changes 508 in future annual temperature for each city, relative to the average baseline temperature 509 510 between 1985–2014. The population size at baseline and projected temperatures were aggregated by grid-level data at the city level. Finally, we calculated the childhood anemia 511 burden attributable to climate change as the proportion of excess childhood anemia cases 512 related to temperature increase in the total numbers of children under 5 years at baseline in 513 each city, multiplying 100,000 person-year. The excess cases were aggregated by country, 514 515 decade, and scenario, and were then used to calculate the future changes in childhood 516 anemia burden due to climate warming. We then obtained total attributable burden of 517 childhood anemia by taking average value of the attributable burden calculated by all 20 GCMs. The empirical confidence intervals (eCIs) were calculated by generating 1,000 518 samples of the coefficients through Monte Carlo simulations, assuming a normal 519 520 distribution for the estimated coefficients, to quantify the uncertainty in estimations of the E-R relationships and the variability in temperature projections for each of the twenty 521

522 GCMs. We finally obtained eCIs corresponding to the 2.5<sup>th</sup> and 97.5<sup>th</sup> percentiles of the 523 distribution of the results across coefficients and 20 GCMs.

All statistical analyses were conducted in the R software (Version 4.0.3, R Project for Statistical Computing) with the "hyfo" package for bias-correction process, and "Im4e" package for mixed effects multivariable logistic models, and the "meta" package for metaregression analyses. All statistical tests were two–sided, and a P–value <0.05 was considered statistically significant.

#### **Supplemental Information:** 529

Supplemental Information includes one figures and fourteen tables. 530

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#### Author contributions: 536

539

540

HK and RC are the joint corresponding authors and contributed to the conceptualization, 537

funding acquisition, project administration and supervision of the study. YZ and CH are 538

the joint first authors and contributed to data curation, investigation, methodology, formal

analysis, and writing-original draft. AG and AV contributed to methodology, formal

analysis and writing-original draft. CL, JB and YC contributed to methodology, software 541

and validation. LZ contributed to data curation. LK contributed to writing-original draft 542

### **Declaration of interests:** 543

The authors declare that they have no conflict of interest. 544

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- 763

- 764 **Figure titles and legends**
- Figure 1. Annual temperature exposure (A) and prevalence of childhood anemia (B)
  for each cluster in sub-Saharan Africa, 2003–2020.
- Figure 2. Odds ratios (95% confidence intervals) of childhood anemia per 1°C increase in annual temperature, classified by different regions or countries of sub-
- 769 Saharan Africa.
- Figure 3. Projected temperature change from the baseline period (1985–2014) under
- 771 different climate change scenarios in sub-Saharan Africa.
- (A) Temporal trends of annual temperature change from the baseline period under three
  climate scenarios from 1985 to 2100. The shaded areas are the interquartile ranges of
  predicted temperatures from twenty general circulation models. (B) Spatial distribution of
  temperature change from the baseline period in city in 2030, 2060, 2090. SSP=Shared
  Socioeconomic Pathway.
- Figure 4. Projected excess cases of childhood anemia per 100,000 person-year
   attributable to climate warming in sub-Saharan Africa under different scenarios.
- (A) Spatial distribution of excess cases of childhood anemia per 100,000 person-year in
  2090s. (B) Projected excess cases of childhood anemia per 100,000 person-year
  attributable to climate warming (red bar) by year, country and scenario. The black vertical
  lines represent the 95% empirical confidence intervals. SSP=Shared Socioeconomic
  Pathway.
- 784

|                              | N       | Percentage<br>(%) | Anemia         | Temperature (°C) |      |      |
|------------------------------|---------|-------------------|----------------|------------------|------|------|
| Region/Country               |         |                   | Prevalence (%) | Mean±SD          | Min  | Max  |
| Eastern Africa               | 103,628 | 37.63%            | 55.36%         | 21.7±2.6         | 12.6 | 32.6 |
| Burundi                      | 16,974  | 6.16%             | 53.31%         | 21.0±1.4         | 17.7 | 26.1 |
| Ethiopia                     | 19,529  | 7.09%             | 55.01%         | 21.3±4.2         | 12.6 | 32.6 |
| Malawi                       | 10,992  | 3.99%             | 65.59%         | 22.2±1.7         | 16.9 | 26.6 |
| Mozambique                   | 4,355   | 1.58%             | 66.25%         | 23.6±1.5         | 18.4 | 26.7 |
| Rwanda                       | 10,636  | 3.86%             | 42.51%         | 19.5±1.7         | 15.6 | 25.7 |
| Tanzania                     | 14,766  | 5.36%             | 56.39%         | 23.4±2.4         | 15.4 | 27.1 |
| Uganda                       | 7,842   | 2.85%             | 58.70%         | 23.0±1.8         | 15.1 | 27.1 |
| Zambia                       | 7,722   | 2.80%             | 58.43%         | 21.9±1.3         | 19.0 | 26.1 |
| Zimbabwe                     | 10,812  | 3.93%             | 51.06%         | 21.3±1.8         | 17.4 | 27.0 |
| Central Africa               | 58,407  | 21.21%            | 63.94%         | 24.1±2.3         | 16.2 | 29.8 |
| Angola                       | 11,198  | 4.07%             | 64.99%         | 22.8±1.9         | 18.6 | 25.9 |
| Cameroon                     | 23,502  | 8.53%             | 62.24%         | 24.7±2.6         | 18.7 | 29.8 |
| Congo Democratic<br>Republic | 20,692  | 7.51%             | 65.32%         | 24.1±2.1         | 16.2 | 27.3 |
| Gabon                        | 3,015   | 1.09%             | 63.81%         | 25.0±0.7         | 22.6 | 26.1 |
| Western Africa               | 103,586 | 37.62%            | 73.42%         | 27.7±1.4         | 19.8 | 31.8 |
| Benin                        | 17,652  | 6.41%             | 67.14%         | 27.4±0.8         | 25.5 | 30.1 |
| Burkina Faso                 | 17,532  | 6.37%             | 89.25%         | 28.8±0.6         | 26.8 | 30.7 |
| Cote d'Ivoire                | 5,560   | 2.02%             | 75.29%         | 26.3±0.6         | 24.2 | 28   |
| Gambia                       | 3,256   | 1.18%             | 53.38%         | 27.7±1.7         | 25.1 | 30.3 |
| Ghana                        | 7,549   | 2.74%             | 74.84%         | 27.0±1.1         | 25.1 | 29.5 |

 Table 1. Summary statistics for the country-specific distribution of children under 5 years

 and annual temperature.

| Liberia         | 2,149   | 0.78%   | 71.20% | 25.9±0.4 | 24.6 | 26.6 |
|-----------------|---------|---------|--------|----------|------|------|
| Nigeria         | 10,168  | 3.69%   | 68.79% | 27.0±1.2 | 19.8 | 29.8 |
| Senegal         | 36,840  | 13.38%  | 71.68% | 28.0±1.7 | 23.7 | 31.8 |
| Togo            | 2,880   | 1.05%   | 71.01% | 27.2±1.0 | 24.8 | 29.4 |
| Southern Africa | 9,756   | 3.54%   | 49.36% | 17.0±4.6 | 12.8 | 25.9 |
| Eswatini        | 3,546   | 1.29%   | 43.99% | 19.1±2.1 | 15.4 | 25.8 |
| Lesotho         | 3,876   | 1.41%   | 51.24% | 16.3±2.8 | 12.8 | 25.7 |
| Namibia         | 1,529   | 0.56%   | 51.21% | 22.4±2.0 | 15.9 | 25.9 |
| South Africa    | 805     | 0.29%   | 60.50% | 19.4±2.0 | 13.8 | 24.1 |
| Total           | 275,377 | 100.00% | 63.76% | 24.4±3.7 | 12.6 | 32.6 |

Abbreviation: SD: Standard Deviation.

A

50°N 40°N 30°N 20°N 10°N 0° 10°S 20°5 30°S 40°S 50°S 30°W 20°W 10°W

10°E 20°E 30°E 40°E 50°E 60°E

0°





| Region/country     | Odds ratio estimates |  |  |  |
|--------------------|----------------------|--|--|--|
| Sub-Saharan Africa | 1.138 (1.134, 1.142) |  |  |  |
| Eastern Africa     | 1.112 (1.103, 1.121) |  |  |  |
| Burundi            | 1.066 (1.041, 1.092) |  |  |  |
| Ethiopia           | 1.032 (1.022, 1.041) |  |  |  |
| Malawi             | 1.152 (1.121, 1.184) |  |  |  |
| Mozambique         | 1.109 (1.057, 1.164) |  |  |  |
| Rwanda             | 1.102 (1.067, 1.139) |  |  |  |
| Tanzania           | 1.124 (1.107, 1.142) |  |  |  |
| Uganda             | 1.150 (1.112, 1.190) |  |  |  |
| Zambia             | 1.047 (1.003, 1.093) |  |  |  |
| Zimbabwe           | 0.997 (0.971, 1.024) |  |  |  |
| Central Africa     | 1.123 (1.112, 1.134) |  |  |  |
| Angola             | 1.085 (1.059, 1.112) |  |  |  |
| Cameroon           | 1.079 (1.062, 1.097) |  |  |  |
| Congo              | 1.164 (1.147, 1.182) |  |  |  |
| Gabon              | 1.225 (1.087, 1.380) |  |  |  |
| Western Africa     | 1.042 (1.029, 1.055) |  |  |  |
| Benin              | 1.120 (1.066, 1.177) |  |  |  |
| Burkina Faso       | 1.058 (0.952, 1.176) |  |  |  |
| Cote d'Ivoire      | 1.020 (0.888, 1.171) |  |  |  |
| Gambia             | 1.212 (1.141, 1.287) |  |  |  |
| Ghana              | 1.008 (0.904, 1.123) |  |  |  |
| Liberia            | 1.293 (0.964, 1.733) |  |  |  |
| Nigeria            | 1.123 (1.081, 1.167) |  |  |  |
| Senegal            | 1.012 (0.994, 1.029) |  |  |  |
| Тодо               | 0.930 (0.831, 1.041) |  |  |  |
| Southern Africa    | 1.010 (0.970, 1.052) |  |  |  |
| Eswatini           | 1.041 (0.992, 1.093) |  |  |  |
| Lesotho            | 0.923 (0.877, 0.971) |  |  |  |
| Namibia            | 0.916 (0.851, 0.986) |  |  |  |
| South Africa       | 1.017 (0.930, 1.112) |  |  |  |
|                    |                      |  |  |  |





