

1 **Association of high ambient temperature with daily hospitalization for cardiorespiratory**  
2 **diseases in Brazil: A national time-series study between 2008 and 2018**

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32

33 **Abstract**

34 Further research is needed to examine the nationwide impact of temperature on health in Brazil, a  
35 region with particular challenges related to climate conditions, environmental characteristics, and  
36 health equity. To address this gap, in this study, we looked at the relationship between high ambient  
37 temperature and hospital admissions for circulatory and respiratory diseases in 5,572 Brazilian  
38 municipalities between 2008 and 2018. We used an extension of the two-stage design with a case  
39 time series to assess this relationship. In the first stage, we applied a distributed lag non-linear  
40 modeling framework to create a cross-basis function. We next applied quasi-Poisson regression  
41 models adjusted by PM<sub>2.5</sub>, O<sub>3</sub>, relative humidity, and time-varying confounders. We estimated  
42 relative risks (RRs) of the association of heat (percentile 99<sup>th</sup>) with hospitalization for circulatory  
43 and respiratory diseases by sex, age group, and Brazilian regions. In the second stage, we applied  
44 meta-analysis with random effects to estimate the national RR. Our study population includes  
45 23,791,093 hospital admissions for cardiorespiratory diseases in Brazil between 2008 and 2018.  
46 Among those, 53.1% are respiratory diseases, and 46.9% are circulatory diseases. The robustness  
47 of the RR and the effect size varied significantly by region, sex, age group, and health outcome.  
48 Overall, our findings suggest that i) respiratory admissions had the highest RR, while circulatory  
49 admissions had inconsistent or null RR in several subgroup analyses; ii) there was a large  
50 difference in the cumulative risk ratio across regions; and iii) overall, women and the elderly  
51 population experienced the greatest impact from heat exposure. The pooled national results for the  
52 whole population (all ages and sex) suggest a relative risk of 1.29 (95% CI: 1.26; 1.32) associated  
53 with respiratory admissions. In contrast, national meta-analysis for circulatory admissions  
54 suggested robust positive associations only for people aged 15-45, 46-65, >65 years old; for men  
55 aged 15-45 years old; and women aged 15-45 and 46-65 years old. Our findings are essential for  
56 the body of scientific evidence that has assisted policymakers to promote health equity and to  
57 create adaptive measures and mitigations.

58

59 **Keywords**

60 Temperature; Heat; Hospitalization; Respiratory diseases; Cardiovascular diseases

61

## 62 1. INTRODUCTION

63 Exposure to high ambient temperature has been linked to increased risk of several adverse  
64 health outcomes, causing significant excess morbidity and mortality (Phung et al., 2016). Between  
65 2000 and 2019, non-optimal ambient temperatures were associated with more than 5 million deaths  
66 per year globally, of which 0.91% of all deaths were related to heat exposure (Zhao et al., 2021).  
67 Cardiorespiratory diseases have often been identified as an important public health concern in the  
68 context of climate-related health effects (Carreras et al., 2015; Ferreira et al., 2019; Jacobson et  
69 al., 2021), given that exposure to extreme temperatures may disturb heart rate (Kaldur et al., 2016),  
70 change blood pressure (Radin et al., 2018), increase respiratory infections (D'Amato et al., 2018),  
71 and contribute to the development of chronic conditions such as chronic obstructive pulmonary  
72 disease – COPD (Tsuji et al., 2019).

73 So far, multiple investigations of the health effects of ambient temperature were essential  
74 for the creation of the initial body of solid scientific evidence that has helped policymakers to  
75 create adaptive measures and mitigations (Tong and Ebi, 2019). However, there is still a gap in  
76 the literature, given that most of the studies were performed in the USA, numerous European  
77 countries, China, and Australia (Rocque et al., 2021; Song et al., 2017). According to a review  
78 study focused on South America (de Sousa et al., 2018), in the year 2018, there were very few  
79 climate-related health effects studies in the continent, which Brazil, the most studied country in  
80 the region, had only 8 investigations. Among those investigations in Brazil, dengue was the main  
81 outcome, and only 2 studies considered cardiorespiratory diseases (de Souza et al., 2012;  
82 Kleinfelder et al., 2009). To update this review study in South America, we searched the literature  
83 for studies related to cardiorespiratory hospital admissions in Brazil after 2018. We found 5  
84 studies, of which two looked at the association between heatwaves and risk of hospitalization in  
85 1,814 Brazilian municipalities (Zhao et al., 2019f, 2019d); one focused on heat exposure and  
86 hospitalization for Chronic Obstructive Pulmonary Disease (COPD) in 1,642 municipalities (Zhao  
87 et al., 2019b); and other two studies focused on the associations between temperature variability  
88 and hospitalization for cardiac arrhythmia (Zhao et al., 2019a) and ischemic heart disease (Zhao  
89 et al., 2019g) in 1,814 Brazilian municipalities. More investigations are still required in South  
90 America, particularly in Brazil, in order to promote health equity under global scenarios suggesting  
91 that extreme temperature events will continue increasing their magnitude, intensity, and frequency  
92 (Fiore et al., 2012; IPCC, 2014; Mahowald, 2011).

93 From a methodological point of view, there is still a lack of studies related to  
94 cardiorespiratory hospital admissions in Brazil that consider: i) a study population from all  
95 Brazilian municipalities (total 5,572), since the Brazilian studies mentioned above have accounted  
96 only for part of the country, less than 50% of the total municipalities; and ii) a modeling framework  
97 that considers covariates related to air quality and humidity, since the literature from other  
98 countries has suggested these variables as important control factors for the association between  
99 health outcomes and weather (Danesh Yazdi et al., 2022; Weinberger et al., 2021), and there is no  
100 study in Brazil to date that accounts for these covariates.

101 Besides the methodological gap, there are the public health issues. First, there are  
102 significant differences in the quality of health/environment and health care among different  
103 population groups in Brazil, which negatively affects health/environment equity. Second,  
104 cardiorespiratory diseases represent a large fraction of the Brazilian burden of diseases. According  
105 to the Ministry of Health in Brazil, in 2020 circulatory and respiratory diseases were responsible  
106 for 357,741 and 148,773 deaths, respectively. By comparison, according to the World Health  
107 Organization (WHO, 2014), in 2012 circulatory and respiratory diseases were responsible for 17.5  
108 and 4 million deaths globally, respectively.

109 Therefore, further research is needed to examine the nationwide impact of temperature on  
110 health in Brazil, a region with particular challenges related to climate conditions, environmental  
111 characteristics, and health equity. This is essential to help policymakers to promote health equity  
112 and create adaptive measures and mitigations. To address this gap, in this study we examined the  
113 association (controlling for spatiotemporal variation, air pollution, and relative humidity) of high  
114 ambient temperature with daily hospital admissions for circulatory and respiratory diseases in  
115 5,572 Brazilian municipalities between 2008 and 2018.

116

## 117 **2. MATERIALS AND METHODS**

### 118 **2.1. Hospital admission data**

119 Hospital admission data were provided by Brazilian Ministry of Health. We accessed  
120 individual hospitalizations nationwide in Brazil at a municipality-level over 11 years (January 1,  
121 2008, and December 31, 2018). Our sample included 129,978,694 hospitalization records (all  
122 cases).

123 Hospital admission data included the date of the event, home municipality, age, sex, and  
124 principal diagnosis according to the International Classification of Diseases, version 10 (ICD-10)  
125 codes. There are 5,572 municipalities in Brazil, which are grouped into five regions: North,  
126 Northeast, Midwest, Southeast, and South (Figure S1).

127 We accounted only for hospitalizations due to respiratory diseases (ICD-10 codes J00-J99)  
128 and circulatory diseases (ICD-10 codes I00-I99). This filter yielded 12,621,970 (9.71% of the total  
129 hospitalizations in Brazil) records of hospitalizations for respiratory diseases and 11,169,123  
130 (8.59% of the total hospitalizations in Brazil) hospitalizations for circulatory diseases.

131 We grouped the hospital admission data by summing the number of hospitalizations for  
132 each municipality and the date between January 1, 2008, and December 31, 2018. We also  
133 stratified the number of hospitalizations by sex and age group (15-45, 46-65, and >65 years). Each  
134 group representing the number of hospitalizations consists of a complete time series of equally  
135 distributed observations in each municipality and on each day of the study period.

136

## 137 **2.2. Temperature data**

138 Temperature data was retrieved from satellite remote sensing, specifically from the  
139 European Centre for Medium-Range Weather Forecasts (ECMWF). The data includes surface  
140 temperature (°C) derived from ERA5-Land, with a spatial resolution of 9km and temporal  
141 resolution of 6 hours. Brazil has a total land area of approximately 8,515,767 km<sup>2</sup>. The smallest  
142 municipality in terms of land area is located in the state of Minas Gerais (Southeast region), with  
143 a land area of 3.5 km<sup>2</sup>, while the largest is located in the state of Pará (North region), with a land  
144 area of 159,533 km<sup>2</sup>. On average, a municipality in Brazil has a land area of approximately 1,526  
145 km<sup>2</sup>, with a standard deviation of 3,465 km<sup>2</sup>. The median land area of a municipality in Brazil is  
146 approximately 487 km<sup>2</sup>.

147 We calculated the daily mean temporal resolution for surface temperature. Finally, we  
148 aggregated the temperature data to the municipality level using a zonal mean, taking into account  
149 the geographic location of the boundaries of each Brazilian municipality.

150

## 151 **2.3. Covariates**

152 We considered two groups of covariates – air pollutants and weather conditions. For air  
153 pollutants, we considered particulate matter with an aerodynamic mass diameter below 2.5 µm

154 (PM<sub>2.5</sub>), measured in µg/m<sup>3</sup>; and tropospheric ozone (O<sub>3</sub>), measured in ppb. Air pollution data were  
155 also obtained from satellite observations by ECMWF. The ECMWF employs the Copernicus  
156 Atmosphere Monitoring Service (CAMS). The CAMS service runs ensemble models using several  
157 satellite observations and emission inventories amongst other predictors.

158 The air pollution data have a spatial resolution of 12.5 km and a temporal resolution of 6  
159 hours. We calculated the temporal resolution for the daily mean and then aggregated the data by  
160 the municipality using a zonal mean. For weather conditions, we considered relative humidity (%),  
161 which had the same source and spatiotemporal resolution as for the temperature data.

162 The validation for the CAMS global model is reported by (Inness et al., 2018). Particularly  
163 for PM<sub>2.5</sub>, they validate the CAMS global model using ground observations from the Aerosol  
164 Robotic Network (AERONET). AERONET has over 500 stations worldwide that measure spectral  
165 Aerosol Optical Depth (AOD) using sun photometers. Approximately 27 AERONET stations are  
166 located in Brazil. The validation estimates the mean bias and standard deviation of the satellite's  
167 instruments (included in the CAMS model for aerosols) relative to AERONET data. In South  
168 America, the bias of the satellite's instruments is slightly smaller, with an approximate bias of -  
169  $0.006 \pm 0.128$ . Another investigation shows that the CAMS estimates in South America have a  
170 root mean square error (RMSE) of 0.268 (compared to AERONET stations) (Gueymard and Yang,  
171 2020). Previous studies have shown that AERONET observation sites in South America have  
172 significant representativity for AOD measured by Moderate Resolution Imaging  
173 Spectroradiometer (MODIS) aboard TERRA and AQUA satellites (Hoelzemann et al., 2009),  
174 which is an instrument included in the CAMS model.

175 The relationship between AOD and PM<sub>2.5</sub> is based on the remote sensing process of  
176 detecting aerosols (fine solid and/or liquid particles suspended in the air). In short, the radiation  
177 interacts with aerosols in the atmosphere resulting in distortion. This distortion is estimated by the  
178 radiative transfer model and can be converted into aerosol loading, defined as AOD (Kumar et al.,  
179 2007). Given that AOD and PM<sub>2.5</sub> are based on the primary source (presence of aerosols in the  
180 atmosphere), there is a strong positive relationship between AOD and PM<sub>2.5</sub>. Several studies have  
181 explored this relationship (Naresh, 2010; Xie et al., 2015; Xu and Zhang, 2020; Yang et al., 2019).  
182 Models predicting fine particle using AOD retrievals and ground-based measurements have  
183 reported good performance across regions worldwide, including a cross-validated R<sup>2</sup> of 0.73 in  
184 Switzerland (de Hoogh et al., 2018), 0.87 in China (Hu et al., 2019), 0.88 in Northeastern USA

185 (Kloog et al., 2014), and 0.82 in Brazil (Gonçalves et al., 2018). The CAMS global model is  
186 considered one of the strongest models used in air pollution and epidemiological studies.

187 Several other studies have used ECMWF data as source for PM<sub>2.5</sub> and meteorology  
188 predictions (temperature and humidity), including investigations in i) Portugal on the association  
189 between prevailing circulation patterns and particles - PM<sub>10</sub> and PM<sub>2.5</sub> (Cavaleiro et al., 2021), ii)  
190 Bavaria, Germany, on the relationship between weather variables and ambulatory visits due to  
191 chronic obstructive pulmonary disease (Ferrari et al., 2012), iii) Africa on potential predictability  
192 of malaria (Tompkins and di Giuseppe, 2015), and iv) in Brazil on the potential for a dengue  
193 epidemic during the 2014 World Cup (Lowe et al., 2014) and on the associations of PM<sub>2.5</sub>-related  
194 wildfire emissions and hospital admissions (Weeberb J Requia et al., 2021), birth weight (Requia  
195 et al., 2022a), preterm birth (Requia et al., 2022b), and birth defects (Weeberb J. Requia et al.,  
196 2021).

197 Regarding the validation of O<sub>3</sub>, it was made by comparing the reanalysis data with surface  
198 measurements from observations from the WMO's Global Atmosphere Watch (GAW) program  
199 (Galbaly et al., 2013). Their findings support that mean CAMS reanalysis O<sub>3</sub> agrees with the  
200 surface data to within 10% for most years and has a smaller bias in the period after 2003.

201

#### 202 **2.4. Statistical analysis**

203 We used an extension of the two-stage design with a case time series (Gasparrini et al.,  
204 2022) to assess the association of low and high ambient temperature with daily hospital admissions  
205 for circulatory and respiratory diseases in Brazil between 2008 and 2018. Specifically, we  
206 performed the first stage by Brazilian region by using municipality-level series. In the second  
207 stage, we estimated the national effects. The statistical analysis in each stage is explained below.

208 In the first stage, we applied a distributed lag non-linear modeling (DLNM) framework  
209 (Gasparrini et al., 2010) to create a cross-basis function. This function represents an association  
210 with ambient temperature, allowing non-linearity and temporal dependencies through the lag  
211 component. We defined the cross-basis function with natural cubic splines for the temperature  
212 exposure-response, and two internal knots (at the 10<sup>th</sup>, 75<sup>th</sup>, and 90<sup>th</sup> temperature percentiles) to  
213 represent non-linearity in the exposure-response relationships. We accounted for the lagged effect  
214 of temperature up to 3 days of lag (0-3 days) by using interger as parameterization function. This  
215 lag was chosen based on a previous study in Brazil on ambient heat and hospitalization for COPD

216 (Zhao et al., 2019c). Although this choice has a support in the literature, we are aware that this is  
217 still a sensitive decision, given that other studies have chosen other lags, including 7 days of lag  
218 and even more. Therefore, in the sensitivity analysis we tested the robustness of our results by  
219 considering other lags (this is detailed below). Finally, given that this is a time series data, in this  
220 cross-basis function, we accounted for a list of municipalities defining groups of observations. All  
221 these parameters for the cross-basis function were chosen based on the Akaike Information  
222 Criterion (AIC).

223 Then, we included the cross-basis function in a generalized conditional quasi-Poisson  
224 regression model, controlling for PM<sub>2.5</sub>, O<sub>3</sub>, relative humidity, and time-varying confounders. The  
225 confounding effect of long-term trends was modeled through smooth function of the year (natural  
226 splines with knots every 3 years). Also, in the quasi-Poisson model, we defined a stratum to control  
227 for temporal trends and seasonality within each municipality. We used a time-stratified sampling  
228 to define strata based on the day of the week, month, calendar year, and municipality of the time  
229 series.

230 The analyses were stratified by sex, age group (15-45, 46-65, and >65 years), health  
231 outcome (respiratory and circulatory hospital admissions), and Brazilian regions (North,  
232 Northeast, Midwest, Southeast, and South). We estimated specific relative risks (RRs) and 95%  
233 confidence interval (CI) for each one of these sub-group analyses. The RRs were derived over 3  
234 days of lag (overall cumulative effect) at extreme percentile defined as 99<sup>th</sup> to represent the extreme  
235 heat associations. Here, we used the temperature of minimum risk of hospitalizations (MRH) as  
236 reference, defined as the risk increment at percentiles 1<sup>st</sup> and 99<sup>th</sup> of the joint temperature  
237 distribution relative to MRH.

238 In the second stage, we applied meta-analysis to estimate the RRs at the national level.  
239 Heterogeneity was examined using the I-squared (I<sup>2</sup>) statistic. A p-value >0.05 and/or I<sup>2</sup> <50% was  
240 considered homogeneous.

241 We used R software, version 4.0.2, including the R packages “dlnm” (Distributed Lag  
242 Non-Linear Models) for the distributed lag modeling framework, “gnm” (Generalized Nonlinear  
243 Models) for the generalized conditional quasi-Poisson regression model, and “mixmeta” for the  
244 meta-analysis.

245

## 246 **2.5. Sensitivity analysis**



247 We performed sensitivity analysis by changing some modeling assumptions. First, we  
 248 adjusted the model only for temporal trends by removing the control for PM<sub>2.5</sub>, O<sub>3</sub>, and relative  
 249 humidity. Then, we tested a model controlling only for PM<sub>2.5</sub>, another model with only O<sub>3</sub>, and  
 250 another one controlling only for humidity. Finally, we tested the sensitivity of our results by  
 251 changing 5, 7, 10, 15, 20, 25, and 30 days of lag.

252

### 253 3. RESULTS

254 Our study population includes 23,791,093 hospitalizations for circulatory and respiratory  
 255 diseases in Brazil between 2008 and 2018. Among those, 12,621,970 (53%) were people  
 256 hospitalized due to respiratory diseases 11,169,123 (47%) due to circulatory diseases. Men were  
 257 the most affected group, representing 52.7% for respiratory admissions and 50.2% for circulatory  
 258 admissions. For the age group, patients aged >65 had the highest proportion, with 25.8%  
 259 hospitalized due to respiratory diseases and 45.1% due to circulatory diseases (Table 1).

260 Table 2 provides the descriptive statistics for temperature and the covariates (relative  
 261 humidity and air pollutants) in Brazil during our study period. Between 2008 and 2018, the mean  
 262 ambient temperature was 23.59°C (sd = 3.79°C), with a minimum of 1.25°C and a maximum of  
 263 34.92°C. Relative humidity had a mean of 77.76% (sd = 13.64%). For air pollutants, we estimated  
 264 a mean concentration of 10.01µg/m<sup>3</sup> (sd = 8.28µg/m<sup>3</sup>) for PM<sub>2.5</sub> and 20.07ppb (sd = 7.01ppb) for  
 265 O<sub>3</sub> (Table 2). Figure 1 shows Brazil's spatial distribution of ambient temperature and hospital  
 266 admissions in the study period.

267

268 Table 1 – Descriptive characteristics of hospital admission events in Brazil, 2008-2018.

Health outcome	Age	Number of hospital admissions (%) <sup>1</sup>		
		Men	Women	All sex
<b>Respiratory hospital admissions</b>	15-45	1,029,372 (8.16)	1,023,403 (8.11)	2,052,775 (16.26)
	46-65	988,490 (7.83)	910,468 (7.21)	1,898,958 (15.04)
	>65	1,630,513 (12.92)	1,632,364 (12.93)	3,262,877 (25.85)
	All ages <sup>2</sup>	6,646,890 (52.66)	5,975,080 (47.34)	12,621,970 (100)
<b>Circulatory hospital admissions</b>	15-45	790,658 (7.08)	1,026,640 (9.19)	1,817,298 (16.27)
	46-65	2,362,171 (21.15)	2,052,674 (18.38)	4,414,845 (39.53)
	>65	2,516,053 (22.53)	2,522,032 (22.58)	5,038,085 (45.11)
	All ages <sup>2</sup>	5,611,256 (50.24)	5,557,867 (49.76)	11,169,123 (100)

269 Notes: <sup>1</sup> the percentages were based on the total number of hospital admissions in Brazil between 2008 and 2018,  
270 which for respiratory hospital admissions were 12,621,970 cases, and for circulatory admissions were 11,169,123  
271 cases. <sup>2</sup> this includes people under 15 years old.

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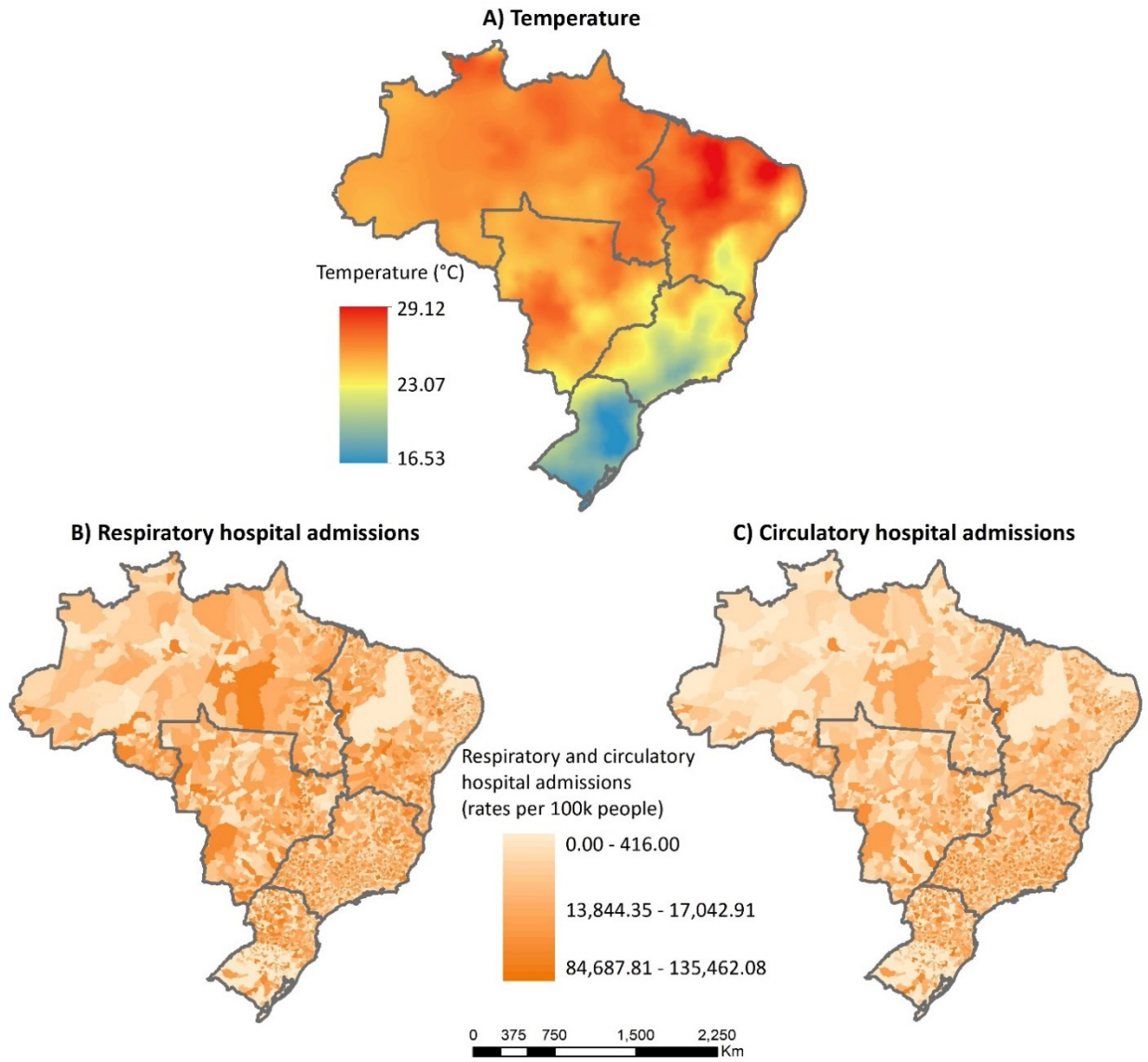
274 Table 2 – Summary statistics for temperature and covariates (humidity and air pollutants) in Brazil, 2008-2018.

<b>Variable</b>	<b>Min</b>	<b>Q1</b>	<b>Mean</b>	<b>SD</b>	<b>Q3</b>	<b>Max</b>
Temperature (°C)	1.25	21.60	23.59	3.79	26.23	34.92
Relative humidity (%)	20.00	69.75	77.76	13.64	88.50	100
PM <sub>2.5</sub> (µg/m <sup>3</sup> )	3.03	5.05	10.01	8.28	11.85	61.35
O <sub>3</sub> (ppb)	3.36	15.22	20.07	7.01	24.77	39.15

275 Note 1: summary statistics were calculated considering the values at the municipality level averaged across the entire  
276 country in the study period.

277 Note 2: Minimum (Min), first quartile (Q1), standard deviation (SD), third quartile (Q3), and maximum (Max).

278



279

280 Figure 1 – Spatial variation of ambient temperature (average between 2008-2018) and cardiorespiratory  
 281 hospital admissions (sum between 2008-2018) in Brazil. Note 1: the polygons in gray represent the  
 282 Brazilian regions. Note 2: we used the total population in each municipality to calculate the rates.

283 Figure 2 shows the overall cumulative effect over 3 days of lag for heat (percentile 99<sup>th</sup>)  
284 exposure to ambient temperature in Brazil (results from meta-analysis). The results indicate robust  
285 positive associations for respiratory hospital admissions in all subgroup analyses during heat  
286 exposure. The pooled national results for the whole population (all ages and sex) suggest a relative  
287 risk of 1.29 (95% CI: 1.26; 1.32) associated with respiratory admissions during heat exposure  
288 (Figure 2). In contrast, national meta-analysis for circulatory admissions suggested robust positive  
289 associations only for people aged 15-45, 46-65, >65 years old; for men aged 15-45 years old; and  
290 women aged 15-45 and 46-65 years old (Figure 2). For circulatory admissions, the pooled national  
291 RR for the whole population was 1.01 (95% CI: 0.99; 1.02). We present in the Supplementary  
292 Materials (Table S2) all the national average relative risk (from the primary analysis, for all sub-  
293 group analyses) along with the estimated heterogeneity test from the meta-analysis.

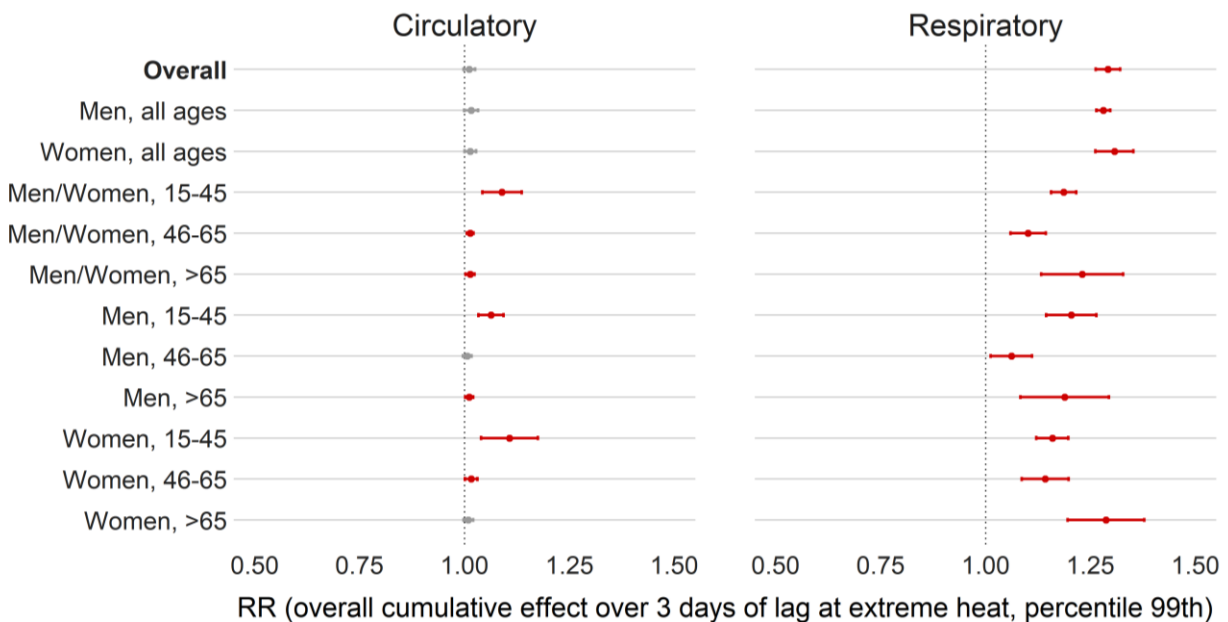
294 There is consistency between the results of the primary analysis and the results of the  
295 sensitivity analysis that considered different covariates in the statistical model (Figure S2).  
296 Overall, the primary coefficients (model adjusted for PM<sub>2.5</sub>, O<sub>3</sub>, humidity, and temporal factors)  
297 associated with cardiorespiratory admissions (Figure S2, panel A) are stable across the different  
298 models, including model adjusted only for PM<sub>2.5</sub> (Figure S2, panel B), model adjusted only for O<sub>3</sub>  
299 (Figure S2, panel C), model adjusted only for humidity (Figure S2, panel D), and model adjusted  
300 only for temporal terms (Figure S2, panel E). In contrast, the results from our primary analysis (3  
301 days of lag) were sensitive to different lags (Figure S3). Overall, we can observe that the higher  
302 lags (5, 7, 10, 15, 20, 25, and 30 days of lag) the higher inconsistency of the coefficients (Figure  
303 S3). Table S3 in Supplementary Materials shows all the RR values at national level from the  
304 primary model and sensitivity analysis (different covariates and different lags for all subgroups  
305 and health outcomes).

306 Figure 3 shows the overall cumulative effect (primary model) over 3 days of lag in each  
307 Brazilian region. The reference of the RR in Figure 3 is based on the minimum risk temperature  
308 (MRT), which is reported in Supplementary Materials (Table S1) with all the coefficients (RR) for  
309 all sub-group analyses and health outcomes. The results suggest a substantial increase in hospital  
310 admission risk for respiratory diseases (all ages and sex) in all regions, with relative risks reaching  
311 1.26 (95% CI: 1.21; 1.31) in the North region, 1.27 (95% CI: 1.24; 1.30) in the Northeast, 1.35  
312 (95% CI: 1.31; 1.40) in the Midwest, 1.30 (95% CI: 1.28; 1.31) in the Southeast, and 1.28 (95%  
313 CI: 1.25; 1.31) in the South region. For circulatory diseases (all ages and sex), we found robust

314 associations only in the South region [1.05 (95% CI: 1.02; 1.07)] and Southeast region [1.02 (95%  
 315 CI: 1.01; 1.03)] (Figure 3). Table S4 in Supplementary Materials shows all the RR values from the  
 316 primary model and sensitivity analysis (different covariates and different lags) for all sub-groups,  
 317 health outcomes, and regions.

318 In Figure S4, we show the bi-dimensional exposure-lag-response association stratified by  
 319 regions for circulatory and respiratory admissions and for the overall analysis (all ages and sex).  
 320 Note that Figure 4 only illustrates a summary of the association between ambient temperature and  
 321 hospital admissions over the lags 1-3 with limited ability to inform inferential information, given  
 322 that the uncertainty is not shown in the 3-D charts. In Figure 4, we show the overall cumulative  
 323 exposure-responses.

324



325

326 Figure 2 - Relative risks (95%CI) heat (percentile 99<sup>th</sup>) exposure in Brazil (results from meta-analysis)  
 327 stratified by health effects, age, and sex.

328 Note 1: This is the overall cumulative effect over 3 days of lag (summing all the contributions up to the maximum  
 329 lag).

330 Note 2: gray color represents the insignificant coefficients (which the RR includes the value 1) and red color represents  
 331 the significant positive associations.

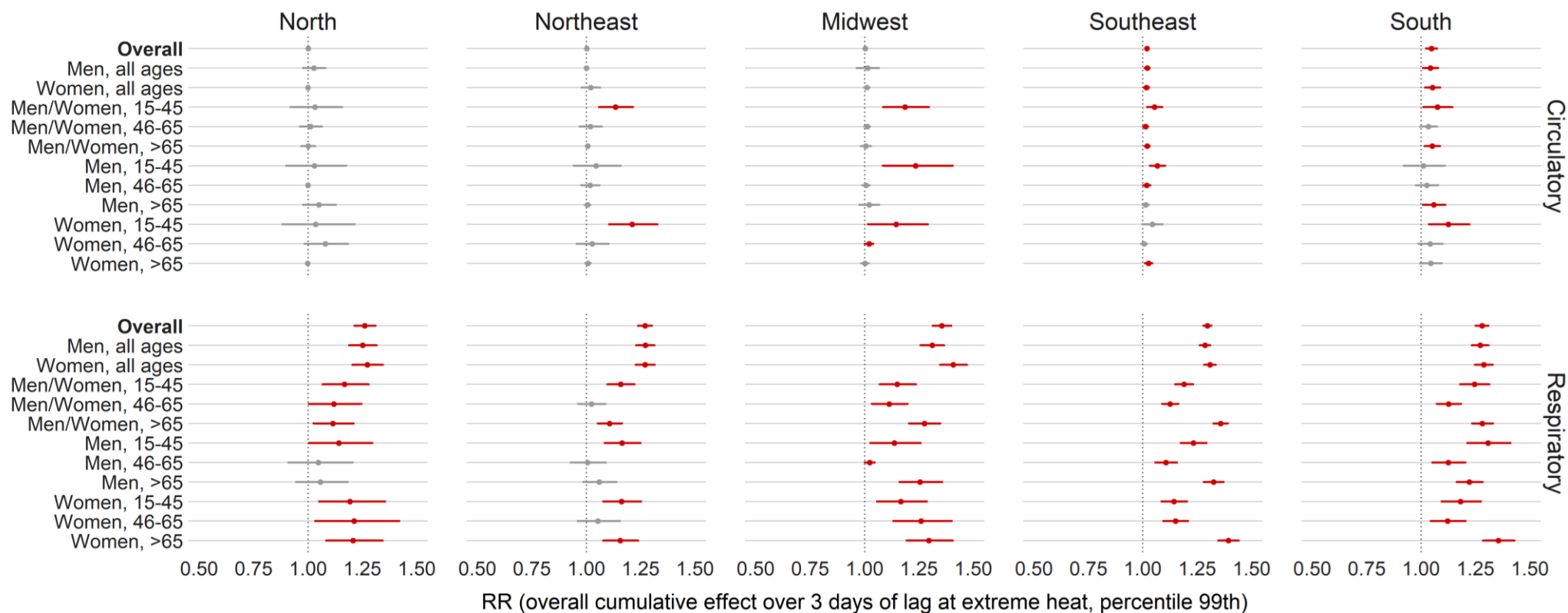


Figure 3 – Relative risks (95%CI) for heat (percentile 99<sup>th</sup>) exposure stratified by regions, health effects, age, and sex.

Note 1: This is the overall cumulative effect over 3 days of lag (summing all the contributions up to the maximum lag).

Note 2: gray color represents the insignificant coefficients (which the RR includes the value 1) and red color represents the significant positive associations.

Note 3: The reference of the RR is based on the minimum risk temperature (MRT), which varied depending on the subgroup analysis. It varied from 23 to 31°C in the North region, 21 to 31°C in the Northeast, 16 to 31°C in the Midwest, 14 to 26.9°C in the Southeast, and 8 to 9.4°C in the South. In Table S1, we show the MRT for each subgroup.

344  
345

### Circulatory

### Respiratory

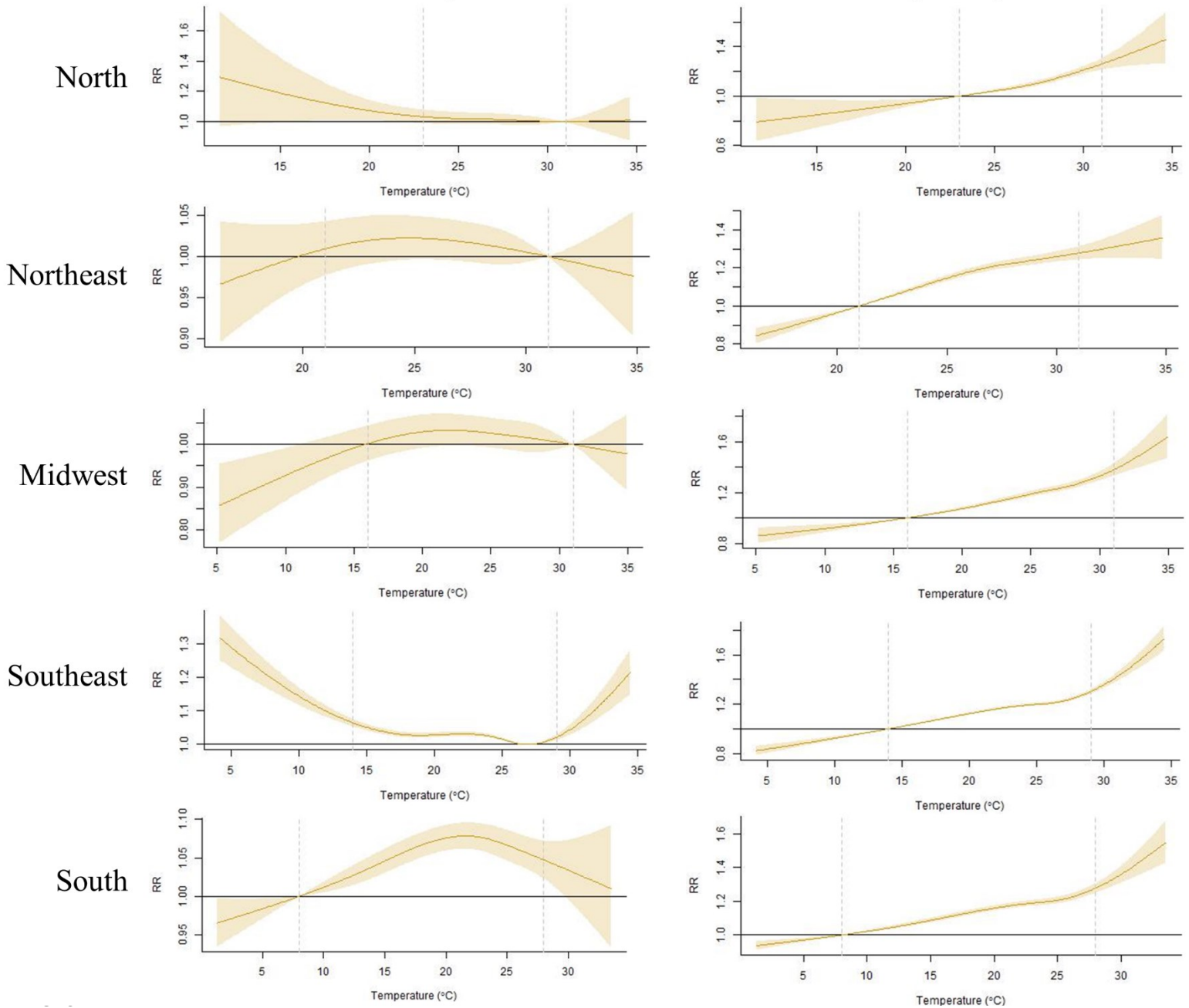


Figure 347- Cumulative risk ratio (95%CI) in ambient temperature for circulatory and respiratory hospital admissions for the overall population. Charts are stratified by health outcome and regions.

Note 349 The reference of the RR is based on the minimum risk temperature (MRT), which for circulatory diseases is 30.8°C in the North 350, 31°C in the Northeast, 31°C in the Midwest, 26.9°C in the Southeast, and 8°C in the South. For respiratory diseases, it is 23°C in the North 351, 21°C in the Northeast, 16°C in the Midwest, 14°C in the Southeast, and 8°C in the South.

Note 352 Vertical lines indicate the 1<sup>st</sup> and 99<sup>th</sup> percentile.

353

354

#### 355 4. DISCUSSION

356 Our results indicate substantial association of high ambient temperature with daily  
357 hospitalization for cardiorespiratory diseases in Brazil. The robustness of the RR and the effect  
358 size varied significantly by region, sex, age group, and health outcome (circulatory and respiratory  
359 admissions). Overall, our findings suggest that i) the highest RR occurred for respiratory  
360 admissions; ii) the RR in several subgroup analyses for circulatory admissions was either null or  
361 inconsistent; iii) there is a large difference in the cumulative risk ratio (curves shown in Figure 4)  
362 between regions; and iv) overall, women and the elderly population were the most impacted by  
363 heat exposure.

364 Our findings are consistent with the literature. In Spain (Iñiguez et al., 2021), the pooled  
365 overall heat effect for circulatory hospitalizations had a null relative risk [in our study it was null  
366 too] and with no statistical significance for respiratory admissions (RR = 1.04, 95% CI: 0.97;1.11)  
367 [in our study it was 1.29 (95% CI: 1.26; 1.32)]. Similar to our study and this study in Spain, which  
368 the risk for circulatory diseases was mostly statistically insignificant, in California, USA, robust  
369 associations related to heat wave was found only for respiratory admissions [1.09 (95% CI: 1.07;  
370 1.12)] (Sherbakov et al., 2018). In contrast, a recent national study in the USA did not find robust  
371 associations for either circulatory or respiratory admissions (Bernstein et al., 2022).

372 Previous studies in Brazil also found similar results. Xu et al. (2020), (Zhao et al., 2019e),  
373 and (Zhao et al., 2019f) looked at the associations between heat exposure and hospitalizations in  
374 Brazil and observed either protective effect or inconsistent associations for circulatory admissions  
375 and robust positive associations for respiratory diseases. They also found a wider confidence  
376 interval with a larger margin of error in the North region. We observed the same issue in our  
377 analysis (Figure 3). We suggest that this is a result of a smaller sample size and a higher variability  
378 in the North region. Another study by Xu et al. (2020) and (Zhao et al., 2019e) found that the  
379 attributable burden was greatest in the Midwest and Northeast. In our study, Midwest had the  
380 highest RR as well [1.35 (95% CI: 1.31; 1.40)], while Northeast and North had the lowest RR,  
381 1.27 (95% CI: 1.24; 1.30) and 1.26 (95% CI: 1.21; 1.31), respectively. The literature has shown  
382 mixed results related to the spatial variability of the risk. Another Brazilian study, for example,  
383 focused on the association between ambient heat and hospitalization for COPD and reported that  
384 the effect was stronger in the Midwest and Southeast regions and minimal in the Northeast (Zhao



385 et al., 2019c). Mixed results were also observed by these Brazilian studies (Xu et al., 2020; Zhao  
386 et al., 2019e, 2019c, 2019f) in the analysis stratified by sex and age groups.

387         This large difference in the risk between regions is explained by the spatial distribution of  
388 the environmental and socioeconomic conditions in Brazil. The Brazilian regions are very  
389 heterogeneous between each other in terms of temperature and also social/economic factors. As  
390 we mentioned in the introduction, there are substantial disparities in the environmental conditions  
391 and quality of public health across Brazilian regions, which negatively impacts health equity and  
392 it is an important determinant of the heat impacts on hospital admissions. Brazil has different types  
393 of biomes (Amazon forest, Cerrado, Atlantic forest, Caatinga, Pampa, and Pantanal) that are  
394 strongly correlated with ambient temperature (LYRA et al. 2016; Salazar et al. 2007). For example,  
395 the Caatinga biome (mainly located in the northeastern region) has become warmer and drier in  
396 the recent years, while the temperature in the South region had a slight variation (Costa et al. 2020;  
397 Da Silva et al. 2019). Regarding sociodemographic conditions, according to the last census,  
398 Northeast is the region with the lowest average income and the Southeast the one with the highest  
399 average income. While the Southeast concentrates the largest urban areas and industries, the  
400 Northeast region has the highest rural population in Brazil, with more than a quarter of the  
401 population living in the countryside.

402         The regional risk observed in our study highlights the importance of considering the spatial  
403 distribution of environmental and socioeconomic conditions when designing policies and  
404 interventions to mitigate the health effects of heat waves. The fact that certain biomes and regions  
405 are more vulnerable to the impacts of high temperature suggests that targeted interventions may  
406 be necessary to address the health equity concerns. Additionally, the disparities in income and  
407 rural-urban divide across regions are also crucial factors that must be taken into account when  
408 designing interventions. Overall, these findings underscore the need for a holistic approach that  
409 considers both the environmental and social determinants of health in order to promote health  
410 equity and reduce the impacts of heat waves in Brazil.

411         Our study has some limitations. First, because of individual perception and decision to seek  
412 medical attention after symptoms of cardiorespiratory disease, some individuals went to the  
413 hospital on the first day of exposure, some others waited until symptoms became too severe and  
414 went to the hospital five days after exposure, some others had acute symptoms but did not go to  
415 the hospital, etc. Because we used a time series framework, this problem may have influenced our

416 findings. Second, there is the possibility of some residual bias, even after adjustment for air  
417 pollution, humidity, spatiotemporal factors, and stratification of the analysis by sex and age. Third,  
418 the predicted temperature, especially the concentration of air pollutants (the main covariate  
419 considered in our analyses) may have introduced some errors in the exposure measurement. In  
420 addition, there is the possibility of exposure misclassification given the coarse spatial resolution  
421 of the temperature and air pollution measurements. Fourth, this is a ecological study which cannot  
422 attribute individual level associations for temperature and hospitalization risk. Also, our study  
423 design does not capture the cause-and-effect relationship between exposure to heat and  
424 cardiorespiratory disease. We only assessed the association of high ambient temperature with daily  
425 hospitalization for cardiorespiratory diseases.

426         Although these limitations exist, our study has some strengths. First, our sample includes  
427 more than 23 million hospitalizations nationwide over 11 years. This sample size provides high  
428 statistical power and enhances the generalizability of our findings to the Brazilian population.  
429 Second, we used a modeling method that flexibly describes associations that have potentially  
430 nonlinear and lagged effects in time series data. Environmental stressors, such as exposure to  
431 temperature, frequently demonstrate time-lagged effects, requiring specific models that account  
432 for the temporal dimension of the exposure-response relationship. Third, we accounted for  
433 spatiotemporal trends in our model by using a time-stratified sample to define strata based on the  
434 day of the week, month, calendar year, and community in the time series. This approach reduces  
435 the effects of confounding factors associated with the spatiotemporal trend by controlling for time-  
436 dependent risk factors within municipalities, including the day of the week, season, and long-term  
437 trends, through matching.

438

## 439 **5. CONCLUSIONS**

440         Our findings suggest that extreme heat is associated with a significant increase in daily  
441 hospital admissions, particularly among the elderly population. These results have important  
442 implications for policymakers, as they can help inform adaptive measures and mitigation strategies  
443 to reduce the adverse effects of climate change on public health in Brazil. Overall, our study  
444 contributes to the growing body of evidence highlighting the urgent need for action to address the  
445 serious climate change challenges facing South America.

446

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